NBP Working Paper No. 261

Formation of inflation expectations in turbulent times

Recent evidence from the European Survey of Professional Forecasters

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Abstract

This paper analyses formation of inflation expectations in the euro area. At the beginning we analyse forecast accuracy of ECB inflation projections relative to private sector forecasts. Then, using the ECB Survey of Professional Forecasters, we estimate a general model integrating two theoretical concepts, i.e. the hybrid model of expectations, including rational and static expectations, and the sticky-information (epidemiological) model. Among determinants of inflation expectations we consider – except backward-looking factors – rational expectations assumption and the effects of the ECB inflation projections. We examine whether ECB inflation projections are still important in expectations' formation once the impact of forward-lookingness of economic agents has been taken into account. We also assess the consistency of implicit (perceived) inflation targets with the ECB inflation target.

Our analysis indicates that recent turbulent times have contributed to changes in expectations' formation in the euro area, as the importance of backward-looking mechanisms has decreased and the importance of the perceived inflation target has increased. We also find that the perceived inflation target has remained broadly consistent with the official ECB inflation target in the medium-term. However, the downward trend of the perceived target signals some risks of de-anchoring of inflation expectations. The importance of ECB inflation projections for medium-term private sector inflation expectations has increased over time, but the magnitude of this effect is rather small. However, SPF inflation forecasts remain consistent with the ECB communication, being ether close to ECB projections or between ECB projections and the inflation target.

JEL: D84, E52, E58.

Keywords: Formation of inflation expectations, survey data, euro area, financial crisis, low inflation.

1. Introduction

Management of inflation expectation has a central role in monetary policy, since expectations are important determinants of actual inflation and their formation influences the effectiveness of monetary policy. Under credible monetary policy economic agents believe that deviations of inflation from the central bank inflation target are temporary, so they set wages and prices in the way consistent with preferences of monetary authorities. The financial crisis and low inflation regime emphasise the need to understand central banks' management of expectations and the impact of monetary policy on private sector expectations. As pointed out by Yellen (2016), "with nominal short-term interest rates at or close to their effective lower bound in many countries, the broader question of how expectations are formed has taken on heightened importance".

Central banks manage expectations by announcing their medium-term inflation targets and communicating with the public. The aim of communication is to increase monetary policy transparency and to reduce economic uncertainty. Qualitative communication refers to formal statements and reports and to more informal speeches and interviews, whereas quantitative communication means publication of central bank macroeconomic projections. In recent years several authors have analysed how inflation projections of central banks affect private sector forecasts (e.g. Hubert, 2015a; Pedersen, 2015, Hattori et al., 2016, Łyziak and Paloviita, 2017). However, in these studies, the simultaneous role of rational expectations assumption has been excluded *a priori*, while the role of official or perceived inflation targets has not been analysed. Our study, which continues our earlier analysis of the euro area inflation expectations (Łyziak and Paloviita, 2017), aims to fill this gap in the literature. We attempt to analyse simultaneously the importance of the rational expectations hypothesis and published ECB inflation projections for the formation of private sector expectations. We also study whether implicit inflation targets perceived by economic agents are consistent with the official ECB inflation target.

Our study uses ECB inflation projections and inflation forecasts in the ECB Survey of Professional Forecasters (ECB SPF). At the beginning we compare forecast accuracy of the ECB projections and SPF forecasts. Then we assess the formation of short- and medium-term SPF inflation expectations by estimating two models of expectations, i.e. the hybrid model, including rational and static expectations, and the sticky-information (epidemiological) model, emphasizing the role of ECB inflation projections for SPF forecasts. We also derive a combined model, integrating these two approaches, in order to assess the relative weights of different factors in expectations' formation.

Our sample period is 1999Q1-2016Q3, which covers recent years characterised by the financial crisis, sovereign debt crisis and low inflation regime. Estimating sub-sample and rolling regressions we examine possible changes of expectations' formation over time.

The paper is organised in the following way. Related literature is summarised in section 2 and the data are described in section 3. The empirical model of expectation formation and its results are presented in section 4. Concluding remarks are provided in section 5.

2. Review of the literature

There are many potential reasons why private sector agents can use central bank forecasts. If the central bank has more resources, better forecasting models or more detailed and timely information about economics developments, it may be able to assess future price developments more accurately than the private sector. In addition, central bank projections may include useful information about economic developments and/or future monetary policy actions.

Our analysis is closely linked to studies testing the impact of central bank forecasts on private sector ones or central banks' forecast performance. Romer and Romer (2000) compare forecast accuracy of the Federal Reserve with forecasts from Blue Chip Economic Indicators. They find that the Federal Reserve has substantially more information about future inflation relative to private sector and private sector forecasters revise their forecasts in response to monetary policy signals of the Federal Reserve. Hubert (2015a) finds that inflation projections of five central banks (Canada, Japan, Sweden, Switzerland and the United Kingdom) affect private sector forecasts, proxied by Consensus Economics survey. However, only in the case of Sweden the central bank's projections outperform private sector forecasts. The same conclusions for Japan is presented by Hattori et al. (2016), who analyse the effects of central bank's inflation forecasts and the adoption of inflation targeting regime on private sector inflation forecasts. Contrary to Ehrmann (2015), who shows that in the low inflation regime inflation expectations are less likely to be anchored and are more sensitive to negative than positive inflation surprises, Hattori et al. (2016) find that the effects of Bank of Japan forecasts on private sector expectations are substantial. Therefore, they argue that targeting inflation from below does not necessarily weaken of the effects of published inflation forecasts. Since forecasting performance of the Bank of Japan is not significantly better than that of private sector, the authors conclude that inflation projections of the central bank include useful complementary information for private agents about future monetary policy actions. Lyziak and Paloviita (2017) examine anchoring of inflation expectations in the euro area. Using the ECB SPF as a proxy for private sector expectations, they show that the ECB inflation projections have recently become more important for short- and medium-term expectations of professional forecasters.

Pedersen (2015) investigates changes (updates) of Chilean private sector inflation expectations using Economic Expectation Survey conducted by the Central Bank of Chile. He finds that the central bank's projections have substantial impact of private sector short-term expectations, while the impact on medium-term expectations is weaker. Hubert (2015c) has used structural VAR models to study the impact of central bank forecast shocks on private sector expectations. He finds that the FOMC forecasts, which include policy signals, affect private sector forecasts, proxied by the US Survey of Professional Forecasters (US SPF). In his another study, Hubert (2015b) also uses structural VAR models to study the impact of ECB inflation projection shocks

on private sector forecasts, which are proxied by the ECB SPF and Consensus Forecasts. His results indicate that private sector expectations are substantially influenced by the ECB projections.

Some authors examine also the impact of central bank inflation projections on dispersion of private sector forecasts. Fujiwara (2005) argues that publishing the Bank of Japan's inflation projections decreased forecasts disagreement among private sector forecasters. Ehrmann et al. (2012) examine projections of central banks in 12 advanced economics and they also find that central bank projections decrease dispersion of private sector forecasts. Kotłowski (2015) argues, however, that in Poland the central bank's release of CPI projections neither influences the cross-sectional dispersion nor the level of expectations of professionals. This can be due to the fact that the central bank inflation target is a dominant factor influencing inflation expectations of professionals (Łyziak, 2013).

The above literature is wide and inspiring. However, attempts to identify the reasons why the private sector follows central bank projections on the basis of forecasting accuracy analysis seem to us problematic. If the private sector believes in information advantages of the central bank and follows central bank communication, then, as a consequence, the forecasting errors committed by the private sector become similar to those of the central bank. In this view the lack of statistical difference between average forecasts errors, interpreted in the literature as a lack of incentives to follow central bank communication by the private sector, results simply from the fact that the credible central bank is able to manage expectations.

Our main contribution to the literature is twofold. First, using a combined model integrating two theoretical approaches we examine backward- and forward-looking determinants of private sector expectations and – in this set-up – the role of implicit inflation target and ECB inflation projections. Our approach enables us to assess whether ECB projections are still important in expectations' formation once the forward-lookingness of economics agents has been taken into account. Second, we provide evidence whether the recent turbulent times have contributed changes in expectations' formation.

3. Data analysis

3.1. Data description

Our data set, which covers the period 1999Q1-2016Q3, includes the euro area HICP inflation rate, ECB SPF inflation forecasts, ECB inflation projections and the ECB inflation target. The ECB SPF has been conducted quarterly since 1999Q1.^{1,2} In every survey round, the survey panel consists of around 50 professional forecasters, who represent both financial and non-financial institutions in the European Union. Forecasters are asked to report, inter alia, their expectations for the euro area HICP inflation rate. Six different forecast horizons are distinguished. In addition to current and next two calendar years, inflation is forecasted in the longer-term horizon. It depends on the quarter, in which the survey is conducted: in the Q1 and Q2 survey rounds it refers to forecasts four calendar years ahead, while in the Q3 and Q4 survey rounds — to forecasts five calendar years ahead.³ Inflation is also forecasted in short- and medium-term fixed horizons, i.e. one year and two years ahead, relative to the month for which the latest official release of the HICP inflation rate is available.⁴

In our study we analyse short- and medium-term expectations since 1999Q1. Before 2014, when inflation expectations for the second calendar year (i.e. the year after the next one) have been surveyed only in quarters Q3 and Q4, we use linear interpolation in order to construct expectations for quarters Q1 and Q2. The ECB SPF is always conducted in the first month of every quarter after the HICP inflation rate for the previous month has been released and its results are published in mid-month of the same quarter. The latest HICP inflation rate is presented to survey participants in the ECB SPF questionnaire.

The ECB inflation projections for the euro area are prepared four times a year. The projections are performed twice a year by the ECB-staff and the Eurosystem National Central Banks in the context of the Eurosystem Staff Broad Macroeconomic Projection Exercise (BMPE). Twice a year these projections are made by the ECB-staff in the context of the ECB-staff Macroeconomic Projection Exercise (MPE). The ECB publishes macroeconomic projections for the current and next two calendar years. In order to emphasize the degree of uncertainty attached to inflation projections, up to March 2013 the ECB published inflation projections in the form of ranges. Therefore, we use midpoints of these ranges until 2013Q1. Since June 2013 we use published midpoints of the ranges for inflation. The ECB projections, which are

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¹ Euro area inflation expectations in the ECB SPF have been analysed in many recent studies. See for example, Conflitti (2012), Rich et al. (2016), Tsenova (2012), Andrade and Le Bihan (2013), Kenny et al. (2014) and Dovern and Kenny (2017).

² Data source: http://www.ecb.europa.eu/stats/prices/indic/forecast/html/index.en.html. The ECB SPF is described in detail in Bowles et al. (2007). See http://www.ecb.europa.eu/stats/pdf/spfquestionnaire.pdf?a9c65f6e4b965a8832693dcb0aebff66 for a survey questionnaire in January 2013 and

http://www.ecb.europa.eu/stats/prices/indic/forecast/shared/files/dataset_documentation_csv.pdf?76c07dc372dffabc3fec09d8 cefbf682 for a description of the ECB SPF data set.

³ For example, in the 2014Q1 and 2014Q2 surveys, the long-term referred to calendar year 2018 and in the 2014Q3 and 2014Q4 surveys to calendar year 2019.

⁴ For example, in the 2015Q1 survey (after release of the HICP inflation rate in December 2014) the forecasters were asked to report their expectations for the HICP inflation rate in December 2015 and in December 2016.

⁵ In our data set, ECB inflation projections for the current and next calendar years are available since 2000Q4 and for the year after next calendar year since 2001Q4.

published in the last month of every quarter, are based on information available until the first month of the same quarter. These cut-off dates reflect relatively long forecasting process in the Eurosystem.

The ECB inflation projections for full calendar years are fully comparable with corresponding ECB SPF inflation forecasts (the same target horizon). For our estimation purposes we construct proxies for one-year-ahead and two-years-ahead ECB inflation projections. One-year-ahead projections are constructed as weighted averages of projections for current and next calendar years (see: Gerlach, 2007 and Dovern et al., 2012). Two-years-ahead projections are assumed to be equal to the ECB projections for the calendar year after the next one. These projections have been published only once a year before 2014 (in last quarters). We use linear interpolation in order to construct observations for quarters Q1-Q3 for those years.⁶

All variables in our data set are shown Figure 1 and basic statistics for the full sample (1999Q1-2016Q3) and the two sub-samples (pre-crisis period: 1999Q1-2008Q3 and post-crisis period 2008Q4-2016Q3) are summarised in Table 1. The two sub-samples are separated by the onset of the financial crisis. We can observe that inflation expectations of professional forecasters have been more stable than the actual HICP inflation rate and that they have been either close to the ECB inflation projections or between ECB inflation projections and the inflation target. It suggests some degree of their anchoring. Short-term inflation expectations have been more volatile than the medium-term ones. In the pre-crisis period short-term projections of the ECB were somewhat higher than those of professionals, but the opposite is true in the case of medium-term expectations. After 2006 both actual and expected inflation have been more volatile than previously. After 2012, higher short-term predictions have been observed for professional forecasters than for the ECB.

3.2. Forecasting performance

Comparing forecasting accuracy of ECB inflation projections and SPF forecasts, we construct forecast errors of the ECB projections in two different ways. First, we analyse forecasting accuracy of ECB projections from the point of view of the private sector, i.e. we consider ECB projections from the previous quarter, available for SPF experts when they report their forecasts. Second, being aware of the fact that the previous quarter projections released by the central bank are inevitably based on outdated information⁸, we use ECB projections published

⁶ Since the ECB does not publish long term inflation projections, they are not included in our analysis.

⁷ Lehman Brothers collapsed in September 2008.

⁸ Due to the relatively long forecasting process in the Eurosystem, the delay between the cut-off date and publication of projections is several weeks and therefore ECB projections available at the time of SPF round do not include timely information about economic developments and ECB communication. Real time information is important in forecasting, however. Especially, if economic developments are very volatile and/or the ECB is active in communication, timely information is crucial in forecasting.

in the current quarter, i.e. the same quarter when the ECB SPF is conducted. This seems a reasonable proxy for the current content of ECB communication concerning future price developments, since the same quarter ECB projections include recent price developments and timely information about economic developments, reflected in other (qualitative) forms of ECB communication (speeches and interviews etc.) available to SPF experts. The information delay between the cut-off date of lagged ECB projections and the conduct of the ECB SPF (deadline to reply) is two months (as an example, see Table 2 for important dates in 2013). Instead, the current quarter ECB projections are published only one month after the survey is conducted.

Inflation forecast errors are expressed as the difference between expected and actual inflation. Forecast errors are examined both in the short- and medium-term forecasting horizons.

To compare accuracy of ECB inflation projections and SPF inflation forecasts, we follow Romer and Romer (2000) and estimate the following equation:

$$\left(e_{t|t+n}^{SPF}\right)^2 - \left(e_{t|t+n}^{CB}\right)^2 = \alpha_n + \varepsilon_t. \tag{1}$$

The left-hand-side variable in the above equation measures the difference between mean squared errors of SPF and ECB forecasts, α_n refers to the constant term and ε to the error term. The null hypothesis of this test is that on average there is no statistical difference in forecasting accuracy, i.e. $\alpha_n = 0$. If the estimated constant is statistically different from zero, the test indicates that there are systematic differences in the forecasting accuracy. A positive value of α_n would mean that central bank inflation projections are systematically more accurate than SPF forecasts.

Test results (Table 3) suggest that in the pre-crisis period the average mean squared error (MSE) for short-term expectations is somewhat lower for professionals, but differences between medium-term forecast accuracies are marginal. In the crisis period smaller errors are observed for the ECB in most of the cases. However, the null hypothesis according to which the forecasting errors are equal to each other, is rejected in none of the cases. Therefore, in a statistical sense, the ECB's forecasting accuracy seems not to be significantly different than that of professional forecasters.

Comparable forecasting accuracy of ECB projections and SPF forecasts suggests that on average they are similar to each other. It can result either from paying attention by SPF experts to ECB projections and adjusting their forecasts to central bank views or from similar schemes of processing available information. Therefore, when modelling expectations' formation of professionals, the role of ECB inflation projections should be empirically checked. It motivates the content of the next section and analytical scheme applied there.

4. How professional forecasters form their inflation expectations?

In this section we first study expectations' formation using two theoretical models and then we continue our analysis using a combined model. The first model, called hybrid, is based on rational expectations and static (simple adaptive) expectations, and the second model, based on sticky-information literature, applies the epidemiological approach (Carroll, 2003, 2006) to test the impact of ECB inflation projections on private sector expectations. Following Fair and Shiller (1990) we combine the two models in order to assess various determinants of inflation expectations. Our analyses reveal whether short- and medium-term private sector inflation expectations are anchored and whether the anchor corresponds to the ECB inflation target.

4.1. Two models of expectations' formation

Our analyses are based on two models of expectations, rooted in the theory and used in empirical literature, both of which include backward- and forward-looking determinants of expectations. The first model (M1) is the hybrid model of expectation formation, similar to models used by Gerberding (2001), Carlson and Valev (2002), Heineman and Ullrich (2006) and Łyziak and Mackiewicz-Łyziak (2014). It combines the models of static (simple adaptive) expectations and rational expectations, i.e.:

$$\pi_{t|t+n}^e = \alpha_0 + \alpha^{BL} \pi_{t-1} + \alpha^{FL} \pi_{t+n} + \varepsilon_t, \tag{2}$$

where $\pi^e_{t|t+n}$ denotes inflation expectations set at time t with the horizon t+n, π_{t-1} denotes the most recent HICP inflation rate, known when the ECB SPF is conducted, while π_{t+n} is the future inflation in the month corresponding to forecasting horizon. Time horizons refer to one year (12 months) and two years (24 months) ahead relative to the month for which the latest HICP inflation rate is available for professionals when they form their own expectations, $n=\{4,8\}$. The model can be reduced to the model of rational (unbiased) expectations if $\alpha_0=0$, $\alpha^{BL}=0$ and $\alpha^{FL}=1$.

The constant term in the equation (3) enables us to calculate so-called implicit (perceived) inflation target, to which a part of economic agents anchor their expectations. We assume that the expected inflation (dependent variable) can be expressed as the weighted average of three factors, i.e. the implicit target (constant over time), past inflation and future inflation, with weights reflecting the shares of economic agents using the above factors in setting their expectations. We assume that the shares of agents using static and rational models of expectation formation are given by the estimated parameters $\hat{\alpha}^{BL}$ and $\hat{\alpha}^{FL}$. Using the estimated

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⁹ Strictly speaking, forecast horizons refer to certain months in the future, not quarters (see section 3.1).

constant term, $\hat{\alpha}_0$, we are able to calculate the parameter α^{IT} , which denotes the share of economic agents whose expectations are anchored to the implicit (perceived) target, π^* :

$$\hat{\alpha}_0 = \alpha^{IT} \pi^* \tag{3}$$

Shares of agents, who form their expectations on the basis of the implicit inflation target, past inflation and future inflation are assumed to sum to one, which implies that:

$$\alpha^{IT} = 1 - \hat{\alpha}^{BL} - \hat{\alpha}^{FL} \tag{4}$$

As a result, the implicit inflation target can be calculated on the basis of estimation results of the hybrid model with the formula:

$$\pi^* = \frac{\hat{\alpha}_0}{1 - \hat{\alpha}^{BL} - \hat{\alpha}^{FL}} \tag{5}$$

A similar approach was used by Kabundi et al. (2015). Analysing inflation expectations in South Africa they derive separate implicit targets for analysts, business and trade unions.

The second model (M2) is based on epidemiological models suggested by Carroll (2003, 2006), rooted in the sticky information paradigm (Mankiw and Reis, 2002). Epidemiological models describe the process of the diffusion of information on the basis of models of disease. They are based on the assumption that agents not experienced in macroeconomic forecasting, like consumers, adjust their inflation predictions to information provided by the media, including forecasts of professional experts. Epidemiological models have been used mainly in modelling consumer inflation expectations (e.g. Nunes, 2009; Łyziak, 2013), but they seem to be also adequate to analyse forecasts of professional economists. Using the epidemiological model, we test how ECB SPF experts process central bank projections when they form their own expectations. We estimate the following equation proposed by Carroll (2003):

$$\pi_{t|t+n}^e = \beta^{BL} \pi_{t-1|t+n-1}^e + \beta^{CB} \pi_{t|t+n}^{CB} + \mu_t, \tag{6}$$

where $\pi_{t|t+n}^{CB}$ denotes the most recent central bank inflation projection available at the time when ECB SPF experts make their own forecasts.

We assume that professional forecasters gradually adjust their views to central bank projections and their forecasts converge to central bank forecasts in the long-run. This imposes the restriction: $\beta^{BL} + \beta^{CB} = 1$. Therefore the final model to be estimated is the following:

$$\pi_{t|t+n}^e = \beta^{BL} \pi_{t-1|t+n-1}^e + (1 - \beta^{BL}) \pi_{t|t+n}^{CB} + \mu_t. \tag{7}$$

4.2. Combined model

Using the hybrid and epidemiological models for expectations' formation, we construct a general model, which combines them. It enables us assess the importance of individual factors (backward-looking factors, unbiased expectations, ECB projections and implicit inflation target) in inflation forecasting. Our approach to assess the information content of models M1 and M2 is inspired by Fair and Shiller (1990). We estimate the following equation:

$$\pi_{t|t+n}^{e} = \omega^{M1} \hat{\pi}_{t|t+n}^{e,M1} + (1 - \omega^{M1}) \hat{\pi}_{t|t+n}^{e,M2} + \vartheta_{t}, \tag{8}$$

where $\hat{\pi}_{t|t+n}^{e,M1}$ stands for fitted values from the model M1, while $\hat{\pi}_{t|t+n}^{e,M2}$ for fitted values from the model M2. The weights of both models are assumed to sum to one.

When analysing information content of different models, Fair and Schiller (1990) used ex-ante forecasts instead of fitted values in the equation (7). In contradiction to forecasts, fitted values from both models can be highly correlated. In order to address this problem, we estimate an equivalent version of the equation (7) that subtracts $\hat{\pi}_{t|t+n}^{e,1}$ from both sides of the equation, i.e.:

$$\pi_{t|t+n}^{e} - \hat{\pi}_{t|t+n}^{e,M1} = (1 - \omega^{M1}) \left(\hat{\pi}_{t|t+n}^{e,2} - \hat{\pi}_{t|t+n}^{e,M1} \right) + \vartheta_{t}. \tag{9}$$

Having the weights of both models, i.e. $\widehat{\omega}^{M1}$ and $\widehat{\omega}^{M2} = 1 - \widehat{\omega}^{M1}$, we are able to assess the importance of all explanatory variables in expectations' formation, i.e. past inflation and lagged forecasts, unbiased expectations as well as the central bank inflation projections and the implicit inflation target. The weights of those factors in the combined model – denoted as κ_t^{BL} , κ_t^{FL} , κ_t^{CB} and κ_t^{IT} , respectively – can be calculated in the following way:

$$\kappa_t^{BL} = \widehat{\omega}^{M1} \widehat{\alpha}^{BL} + (1 - \widehat{\omega}^{M1}) \widehat{\beta}^{BL}, \tag{10}$$

$$\kappa_t^{FL} = \widehat{\omega}^{M1} \widehat{\alpha}^{FL},\tag{11}$$

$$\kappa_t^{CB} = (1 - \widehat{\omega}^{M1}) (1 - \widehat{\beta}^{BL}), \tag{12}$$

$$\kappa_t^{IT} = \widehat{\omega}^{M1} (1 - \widehat{\alpha}^{BL} - \widehat{\alpha}^{FL}). \tag{13}$$

4.3. Estimation results

In our estimations we need to use information available for professionals at the time of forecasting. When the ECB SPF is surveyed in the 1st month of every quarter, the same quarter ECB projections are not available for professionals – they are published in the last month of the same quarter. Therefore, we use lagged ECB projections in our estimations, which are summarised in Tables 4-6 and Figures 2-5.

Table 4 presents estimation results in the whole sample period and two sub-samples, separated by the onset of the financial crisis. According to estimated coefficients of the model M1, private sector agents seem to be strongly backward-looking in their expectations' formation – only 11% of agents make unbiased short-term forecasts in the whole sample period. When considering two sub-periods separately, the share of forward-looking agents is lower in the post-crisis period than in pre-crisis period. As far as the formation of medium-term forecasts is considered, the forward-looking element can be ignored independently of the sample period.

The largest group of ECB SPF experts use the implicit inflation target as the focal benchmark for their forecasts. In the formation of short-term inflation forecasts, this share accounts for 65% of all forecasters – it has slightly increased from 64% in the pre-crisis period to 70% in post-crisis period. Medium-term forecasts are more anchored than the short-term ones – 90% of ECB SPF forecasters base their predictions on the perceived target. Interestingly, in the financial crisis period the implicit (perceived) inflation target for short-term expectations has decreased from 1.6% to 1.4% and for medium-term inflation expectations from 1.8% to 1.7%.

The epidemiological model (M2) allows for another source of potentially forward-looking information, i.e. ECB inflation projections. Estimation results (Table 4) suggest that professional forecasters are persistent in their expectations' formation, being reluctant to change their forecasts considerably, especially in the case of the medium-term horizon. The share of experts updating immediately their expectations on the basis of ECB projections is about 20% in the case of short-term forecasts and 10% in the case of medium-term forecasts. Interestingly, in the case of medium-term forecasts, the importance of ECB projections has increased in the post-crisis period compared with the pre-crisis one, but this effect is on the edge of statistical significance.

All in all, estimation results in Table 4 suggests that inflation expectations of professional forecasters in the euro area are formed with a small amount of forward-looking information, but they are also weakly sensitive to changes in current inflation. The dominant mechanisms of expectations' formation, seen through the lenses of the estimated models, rely either on implicit targets, to which experts anchor their expectations or on their previous forecasts, only gradually adjusted towards ECB inflation projections. Since the beginning of financial crisis the fully forward-looking element (given by unbiased forecasts) in the formation of short-term forecasts

has disappeared, while the importance of current inflation has diminished. In the case of medium-term inflation expectations, the willingness of ECB SPF forecasters to adjust their forecasts to the ECB inflation projections has recently increased. On the other hand, both short-and medium-term implicit targets perceived by SPF experts have decreased somewhat recently, deviating more from the ECB official target than in the past.

Our estimations confirm the observation of Lyziak and Paloviita (2017) that even if ECB SPF forecasts have become slightly de-anchored with respect to the ECB inflation target, there has been the compensating effect in terms of increased role of ECB inflation projections for medium-term inflation expectations. It should be noted, however, that the above conclusions are based on two separate models, without assessing their relative usefulness in fitting ECB SPF forecasts.

Therefore the next step of our analysis is to assess the weights of both estimated models, minimizing the squared deviations of fitted values from the combined model (7) from actual ones. Estimation results (Table 5) indicate that in the whole sample period the weight of model M1 (based on rational and static expectations) is slightly lower than the weight of model M2 (the epidemiological model). However, the model of expectations' formation by ECB SPF experts has changed considerably over time. It seems that in the pre-crisis period the epidemiological model is sufficient to adequately reflect inflation forecasts of ECB SPF experts, while in the post-crisis period its weight is reduced to 0.16 for short-term forecasts and to 0.30 for medium-term forecasts.

Combining the importance of single factors affecting expectations in each of two estimated models (Table 4) with derived weights of both models (Table 5) we obtain the final ranking of factors affecting expectations, in line with the equations (9)-(12). The results are presented in Table 6. They indicate that backward-looking factors dominate the formation of private sector forecasts in the whole sample period and, even more strongly, in the pre-crisis period. An important factor affecting ECB SPF experts' expectations is the implicit (perceived) inflation target, the role of which has strongly increased in the recent sub-sample, being almost twice larger than the role of backward-looking factors. As far as the model of unbiased expectations is concerned, we can find its minor importance only for short-term inflation expectations and in the whole sample period. In our combined model including separate forward-looking factor (rational expectations assumption), the ECB inflation projections seem not to be very relevant in the formation of ECB SPF forecasts, although we can observe a slight increase in their role in the case of medium-term SPF forecasts.

The above observations suggest that in recent turbulent years the formation of inflation expectations by professional forecasters has substantially changed. Nowadays those forecasts are driven mainly by the implicit inflation target, which – as discussed above – has been reduced in the recent period, deviating slightly from the ECB official target. It can signal increasing risks of inflation expectations' de-anchoring. Taking into account the combined model of expectation formation we can observe a small compensating effect in terms of the increase of the role of ECB inflation projections in affecting medium-term ECB SPF forecasts.

In order to analyse possible changes in expectations formation in more detail, we repeat the analyses using rolling regressions. In our estimations, the size of the rolling window is 33 quarters: the sample period in the first rolling regression is 1999Q1-2007Q1, in the last one it corresponds to the financial crisis period (2008Q3-2016Q3).

Rolling regression results for the hybrid model M1 (Figure 2) are broadly consistent with findings presented above. The share of forward-looking agents, making unbiased forecasts of future inflation, is low in the case of short-term forecasts and negligible in the case of medium-term forecasts. We can notice downward trend in the importance of this determinant of short-term ECB SPF inflation forecasts. The role of past inflation, even if gradually rising, is not prominent. Majority of professional forecasters base their forecasts, especially medium-term ones, on the perceived inflation target. We can observe that implicit targets are rather stable over time, although at the end of the sample period we can observe a slight downward trend.

Rolling regression results based on the epidemiological model M2 (Figure 3) suggest that ECB SPF experts are rather reluctant to change their forecasts, although the share of experts updating their forecasts on the basis of ECB inflation projections seem to be slightly going up since 2008, especially in the case of short-term forecasts.

As far as the combined model is concerned, it seems that the weights of both models of expectation formation were changing substantially over time (Figure 4). Before the financial crisis the epidemiological model was dominant in both forecast horizons. From 2009 to 2014 the weights of both models were relatively stable and close to each other in the case of medium-term forecasts. Since 2015 the hybrid model based on rational and static expectations has been getting importance again – its weight at the end of the sample period is approximately 70% both for short-term forecasts and medium-term forecasts.

The combined model suggests that private sector expectations' formation has changed since the beginning of the financial crisis (Figure 5). The most significant change is the reduction of backward-lookingness of ECB SPF forecasts and increased role of perceived inflation target. The importance of ECB inflation projections for setting short-term expectations by ECB SPF experts was increasing in 2009-2014, although since mid-2014 this trend has reversed. The degree of rationality of expectations (unbiased condition) has slightly increased after the crisis, but rather temporarily and in the case of short-term forecasts only.

Our empirical analyses suggest that recent years have contributed to substantial changes in expectations' formation in the euro area. The degree of expectations' backward-lookingness has recently decreased, while the importance of perceived target has increased. The perceived inflation target is broadly consistent with the official ECB inflation target in the medium-term; however, we find some evidence of its recent downward trend indicating increasing risks of deanchoring inflation expectations. The fully forward-looking element in the formation of inflation expectations, given by unbiased predictions, is small in the case of short-term forecasts, while null in the case of medium-term ones. According to our combined model, which includes separate forward-looking factor (rational expectations assumption), the direct impact of ECB inflation projections on private sector expectations is quite limited.

5. Conclusions

Recent turbulent times have raised the question whether formation of inflation expectations has changed over time. In such times expectations' formation is especially crucial issue for central banks, who attempt to manage expectations by – among other things – publishing macroeconomic projections. Many authors have studied the impact of central forecasts on private sector inflation forecasting. However, in these studies economic agents have been assumed either backward-looking or to base their expectations on central bank communication. Typically, central bank inflation projections have been treated as the only forward-looking factor affecting expectations. Using our novel approach, we analyse simultaneously the importance of the rational expectations hypothesis and ECB inflation projections for the formation of private sector expectations.

The impact of central bank forecasts on private sector forecasting has been often assessed comparing forecasting performances. It has been argued that only smaller forecast errors of the central bank reveal motivation of the private sector follow central bank communication. Our interpretation is different. We find only marginal differences in forecast errors of the ECB and professionals, which according to our interpretation indicates that the ECB is able to manage expectations – SPF forecasters seem to base their forecasts on the views of the central bank or transform available information in their inflation forecasts in the way similar to the monetary authorities.

To assess the direct role of ECB projections in the formation of SPF inflation forecasts, we apply the model, in which we try to control for different forward-looking and backward-looking determinants of inflation expectations. We find that in recent turbulent times the importance of backward-looking mechanisms has decreased and the importance of the perceived target has increased in expectations' formation. The perceived inflation target has remained broadly consistent with the official ECB inflation target in the medium-term, although we find some signals of its downward trend, reflecting increasing risks of de-anchoring. The immediate impact of ECB inflation projections on private sector inflation forecasts is relatively low once the impact of forward-lookingness of economic agents has been taken into account. However, SPF inflation forecasts remain consistent with the ECB communication – they have been more stable than the actual HICP inflation rate and either close to the ECB inflation projections or between ECB inflation projections and the ECB inflation target.

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Tables and figures

Table 1. HICP inflation rate, ECB SPF inflation expectations and ECB projections

| | HICP | SPF_1Y | SPF_2Y | ECB_1Y | ECB_2Y | | | | | |
|-----------------------------|--------|-------------------|-----------------|--------|--------|--|--|--|--|--|
| Whole sample, 1999Q1-2016Q3 | | | | | | | | | | |
| Mean | 1.738 | 1.629 | 1.749 | 1.661 | 1.612 | | | | | |
| Maximum | 3.800 | 2.401 | 2.109 | 2.900 | 2.000 | | | | | |
| Minimum | -0.400 | 0.755 | 1.221 | 0.375 | 1.300 | | | | | |
| Std. dev. | 0.973 | 0.322 | 0.167 | 0.578 | 0.184 | | | | | |
| Observations | 71 | 71 | 71 | 64 | 60 | | | | | |
| | | Pre-crisis period | , 1999Q1-2008Q3 | | | | | | | |
| Mean | 2.185 | 1.791 | 1.825 | 2.027 | 1.730 | | | | | |
| Maximum | 3.800 | 2.401 | 2.109 | 2.900 | 2.000 | | | | | |
| Minimum | 0.800 | 1.154 | 1.501 | 1.600 | 1.500 | | | | | |
| Std. dev. | 0.602 | 0.224 | 0.109 | 0.329 | 0.155 | | | | | |
| Observations | 39 | 39 | 39 | 32 | 28 | | | | | |
| | | Crisis period, 2 | 2008Q4-2016Q3 | | | | | | | |
| Mean | 1.194 | 1.432 | 1.657 | 1.296 | 1.508 | | | | | |
| Maximum | 2.900 | 1.946 | 1.981 | 2.200 | 1.800 | | | | | |
| Minimum | -0.400 | 0.755 | 1.221 | 0.375 | 1.300 | | | | | |
| Std. dev. | 1.065 | 0.314 | 0.179 | 0.543 | 0.140 | | | | | |
| Observations | 32 | 32 | 32 | 32 | 32 | | | | | |

Note: HICP denotes actual inflation. 1Y (2Y) refers to expectations one year ahead (two years ahead). Inflation expectations are dated at the time when the ECB projections and ECB SPF were published.

Source: own calculations.

Table 2. Timing of ECB projections and ECB SPF forecasts in 2013Q1-2013Q4

| | Previous quarte | r ECB projections | Current quarter | ECB projections | Current quarter ECB SPF | | |
|--------|--|-------------------|-----------------|------------------|-------------------------|------------------|--|
| | Cut-off date Publication date | | Cut-off date | Publication date | Deadline to reply | Publication date | |
| 2013Q1 | 2013Q2 22-Feb-2013 4-March-2013 2013Q3 22-May-2013 3-June 3-2013 | | 22-Feb-2013 | 4-March-2013 | 22-Jan-2013 | 14-Feb-2013 | |
| 2013Q2 | | | 22-May-2013 | 3-June-2013 | 19-Apr-2013 | 09-May-2013 | |
| 2013Q3 | | | 23-Aug-2013 | 2-Sep-2013 | 19-Jul-2013 | 08-Aug-2013 | |
| 2013Q4 | | | 22-Nov-2013 | 2-Dec-2013 | 22-Oct_2013 | 14-Nov-2013 | |

Source: based on ECB documents.

Table 3. Forecasting accuracy of ECB and SPF experts – Romer and Romer (2000) test

| | | whole sample | | pre-crisis crisis sub-sample sub-sample | | | | | | |
|---------------------------------|----------------------------------|-----------------|-----------------|--|-------|-----------------|-------------|-------|-----------------|--|
| | MSE average | | | ge MSE average | | | MSE average | | | |
| | SPF | ECB | <i>p</i> -value | SPF | ECB | <i>p</i> -value | SPF | ECB | <i>p</i> -value | |
| | Previous quarter ECB projections | | | | | | | | | |
| 1-year horizon | 0.905 | 1.032 | 0.464 | 0.789 | 0.952 | 0.310 | 1.034 | 1.125 | 0.748 | |
| 2-year horizon | 1.217 | 1.077 | 0.281 | 0.998 | 1.000 | 0.982 | 1.462 | 1.164 | 0.181 | |
| Current quarter ECB projections | | | | | | | | | | |
| 1-year horizon | 0.887 | 1.026 | 0.451 | 0.769 | 1.227 | 0.214 | 1.027 | 0.787 | 0.269 | |
| 2-year horizon | 1.180 | 1.077 | 0.505 | 0.966 | 0.965 | 0.996 | 1.441 | 1.214 | 0.308 | |

Note: Forecast errors refer to expected minus actual inflation. With previous (current) quarter ECB projections the sample starts in 2001Q1 (2000Q4) for 1-year-ahead forecasts and in 2002Q1 (2001Q4) for 2-years-ahead forecasts. Ordinary least squares with Newey-West HAC standard errors. The *p*-value is for the test of null hypothesis that the private sector errors and ECB errors are equal.

Source: own calculations.

Table 4. Models of expectation formation

| | | SPF 1Y SPF 2Y | | | | | | | | | |
|---------------------|-------------------|--------------------------|----------------------|--|-----------------|--------------------------|----------------------|---|--|--|--|
| | whole sample | pre-crisis sub-sample | crisis sub-sample | F [F-prob] H ₀ : equal coefficients in sub- samples | whole sample | pre-crisis sub-sample | crisis sub-sample | F [F-prob] H ₀ : equal coefficients in subsamples | | | |
| | Model 1 – eq. (3) | | | | | | | | | | |
| Constant | 1.009*** | 1.026*** | 1.040*** | 0.001 | 1.537*** | 1.676*** | 1.524*** | 0.978 | | | |
| term | (0.059) | (0.376) | (0.064) | [0.970] | (0.036) | (0.124) | (0.051) | [0.323] | | | |
| Past | 0.248*** | 0.212** | 0.280*** | 0.371 | 0.123*** | 0.077* | 0.118*** | 0.622 | | | |
| inflation | (0.024) | (0.104) | (0.033) | [0.545] | (0.013) | (0.042) | (0.019) | [0.434] | | | |
| Future | 0.107*** | 0.149* | 0.022 | 2.679 | 0.005 | 0.000 | 0.000 | 0.000 | | | |
| inflation | (0.030) | (0.076) | (0.042) | [0.107] | (0.001) | (0.042) | (0.019) | [0.996] | | | |
| Implicit target | 1.56 | 1.61 | 1.44 | 0.401 [0.529] | 1.76 | 1.82 | 1.73 | 6.861 [0.012] | | | |
| Adj. R ² | 0.73 | 0. | 73 | Х | 0.60 | 0.63 | | x | | | |
| J-prob. | 0.33 | 0. | 69 | x | 0.10 | 0. | 14 | X | | | |
| | | | N | Model 2 – eq. (8) |) | | | | | | |
| Past SPF | 0.808*** | 0.819*** | 0.801*** | | 0.895*** | 0.994*** | 0.859*** | | | | |
| forecast | (0.056) | (0.094) | (0.070) | 0.022 | (0.040) | (0.077) | (0.047) | 2.234 | | | |
| ECB projection | 0.192 | 0.181 | 0.199 | [0.883] | 0.105 | 0.006 | 0.141 | [0.140] | | | |
| Adj. R ² | 0.82 | 0. | 82 | Х | | 0.81 | | X | | | |

Notes: 1Y (2Y) refers to forecasts one year ahead (two years ahead). Actual future inflation is used in the model 1 as a measure of rational expectations. As a consequence, the error term of the estimated equation includes the expectational error of rational expectations (Fair, 1993). Therefore, the two-stage least squares method (2SLS) is used to estimate the model 1. The set of instruments contains three lags of HICP inflation, GDP growth, real interest rate, USD/EUR exchange rate and oil prices. Numbers in parentheses below estimated coefficients are standard errors. *** denotes confidence level at 99 per cent; ** denotes confidence level at 95 per cent; * denotes confidence level at 90 per cent.

Source: own calculations.

Table 5. Relative weights of models of expectation formation

| | | SPF | 1Y | | SPF 2Y | | | |
|---------------------|---------------------|--------------------------|----------------------|---|---------------------|--------------------------|----------------------|---|
| | whole sample | pre-crisis sub-sample | crisis sub-sample | F [F-prob] H ₀ : equal weights in sub-samples | whole sample | pre-crisis sub-sample | crisis sub-sample | F [F-prob] H ₀ : equal weights in sub-samples |
| Model 1 | 0.424*** (0.085) | 0.099 (0.118) | 0.841*** (0.189) | 8.030 | 0.465*** (0.106) | 0.170 (0.142) | 0.696*** (0.132) | 6.461 |
| Model 2 | 0.576 | 1 | 0.159 | [0.006] | 0.535 | 1 | 0.304 | [0.014] |
| Adj. R ² | 0.852 | 0.498 | | X | 0.770 | 0.314 | | х |

Notes: 1Y (2Y) refers to forecasts one year ahead (two years ahead). Numbers in parentheses below estimated coefficients are standard errors. *** denotes confidence level at 99 per cent; ** denotes confidence level at 95 per cent; * denotes confidence level at 90 per cent. *** denotes significance level at 95 per cent; * denotes significance level at 95 per cent; * denotes significance level 90 per cent.

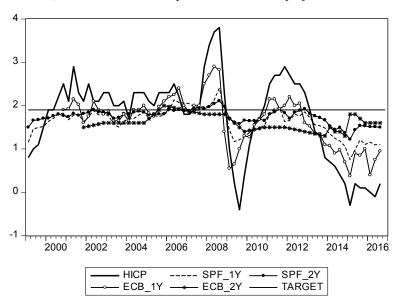
Source: own calculations.

Table 6. Weights of factors affecting SPF forecasts in the combined model of expectation formation (%)

| | | SPF 1Y | | SPF 2Y | | | |
|--|--------------|--------------------------|----------------------|--------------|--------------------------|----------------------|--|
| | whole sample | pre-crisis sub-sample | crisis sub-sample | whole sample | pre-crisis sub-sample | crisis sub-sample | |
| backward-looking factors | 57.1 | 81.9 | 36.3 | 53.6 | 99.4 | 34.3 | |
| unbiased prediction of future inflation | 4.5 | 0.0 | 0.0 | 0.0 | 0.0 | 0.0 | |
| ECB projections | 11.1 | 18.1 | 3.2 | 5.6 | 0.6 | 4.3 | |
| implicit (perceived) inflation target | 27.3 | 0.0 | 60.6 | 40.8 | 0.0 | 61.4 | |

Source: own calculations.

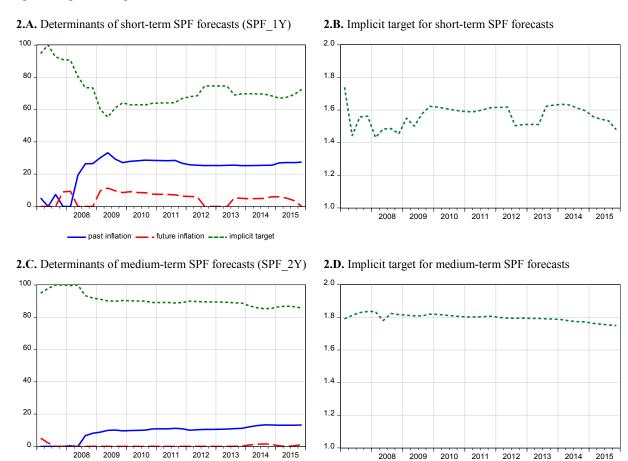
Figure 1. HICP inflation rate, ECB SPF inflation expectations and ECB projections



Note: HICP denotes actual inflation and TARGET the ECB inflation target. 1Y (2Y) refers to expectations one year ahead (two years ahead). Inflation expectations are dated at the time when the ECB projections and ECB SPF were published.

Source: Eurostat, ECB and own calculations.

Figure 2. Weights of factors affecting SPF forecasts in the hybrid model of expectation formation, eq. (1) (%) and implicit targets for expectations



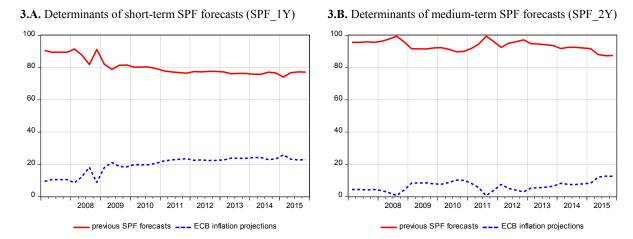
Notes: The weight equals zero if the estimated coefficient is not statistically significant.

___ - future inflation ___- implicit target

Source: own calculations.

past inflation .

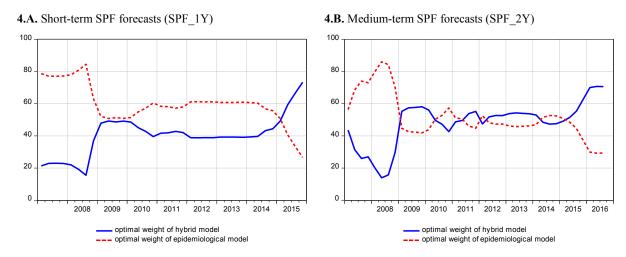
Figure 3. Weights of factors affecting SPF forecasts in the epidemiological model of expectation formation, eq. (6) (%)



Notes: The weight equals zero if the estimated coefficient is not statistically significant.

Source: own calculations.

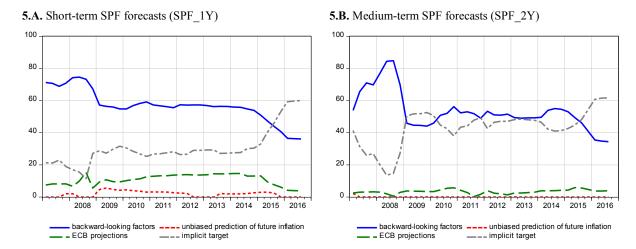
Figure 4. Optimal weights of models of expectation formation (%)



Notes: The weight equals zero if the estimated coefficient is not statistically significant.

Source: own calculations.

Figure 5. Weights of factors affecting SPF forecasts in the combined model of expectation formation (%)



Notes: The weight equals zero if the estimated coefficient is not statistically significant.

Source: own calculations.

